

# Advanced Monte Carlo and Optimization Methods for Optimal Stopping Problems: Part IV

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# Optimal Stopping Problems

Consider the following discrete time **optimal stopping problem**:

$$Y_0^* = \sup_{\tau \in \{0, \dots, \mathcal{T}\}} E[Z_\tau],$$

where

- $(Z_j)_{j \geq 0}$  is an adapted process on a probability space  $(\Omega, (\mathcal{F}_j)_{j \geq 0}, P)$
- $\tau$  is a stopping time with values in  $\{1, \dots, \mathcal{T}\}$ , i.e.  $\{\tau = j\} \in \mathcal{F}_j$

## Question

*How to approximate  $Y_0^*$  in the case when the expectation  $E[Z_\tau]$  cannot be computed in a closed form ?*

# Dual upper bounds

Consider a discrete martingale  $(M_j)_{j=0,\dots,\mathcal{J}}$  with  $M_0 = 0$  w. r. t. the filtration  $(\mathcal{F}_j)_{j=0,\dots,\mathcal{J}}$ . Observe that

$$Y_0^* = \sup_{\tau \in \{0,\dots,\mathcal{J}\}} E^{\mathcal{F}_0} [Z_\tau - M_\tau] \leq E^{\mathcal{F}_0} \max_{0 \leq j \leq \mathcal{J}} [Z_j - M_j] .$$

Hence the r.h.s. with an arbitrary martingale gives an **upper bound** for the Bermudan price  $Y_0$ .

## Question

*What martingale does lead to equality ?*

# Dual upper bounds

## Theorem (Rogers (2001), Haugh & Kogan (2001))

Let  $M^*$  be the (unique) Doob-Meyer martingale part of  $(Y_j^*)_{0 \leq j \leq \mathcal{J}}$ , i.e.  $M_j^*$  is an  $(\mathcal{F}_j)$ -martingale which satisfies

$$Y_j^* = Y_0^* + M_j^* - A_j^*, \quad j = 0, \dots, \mathcal{J}$$

with  $M_0^* := A_0^* := 0$ , where  $A_j^*$  is increasing process which  $\mathcal{F}_{j-1}$ -measurable. Then

$$Y_0^* = \max_{0 \leq j \leq \mathcal{J}} [Z_j - M_j^*], \quad a.s.$$

# Riesz upper bounds

Doob-Meyer decomposition:

$$Y_j^* = Y_0^* + M_j^* - A_j^*, \quad j = 0, \dots, \mathcal{J},$$

and  $Y_{\mathcal{J}}^* = Z_{\mathcal{J}}$  imply **Riesz decomposition**:

$$Y_j^* = E^{\mathcal{F}_j} Z_{\mathcal{J}} + E^{\mathcal{F}_j} (A_{\mathcal{J}}^* - A_j^*)$$

Since  $A_{i+1}^* - A_i^* = Y_i^* - E^{\mathcal{F}_i} Y_{i+1}^* = [Z_i - E^{\mathcal{F}_i} Y_{i+1}^*]^+$ , we get

$$Y_j^* = E^{\mathcal{F}_j} Z_{\mathcal{J}} + E^{\mathcal{F}_j} \sum_{i=j}^{\mathcal{J}-1} [Z_i - E^{\mathcal{F}_i} Y_{i+1}^*]^+.$$

# Riesz upper bounds

## Theorem

If  $Y_i$  is a lower approximation for  $Y_i^*$ , i.e.,  $Y_i \leq Y_i^*$  a.s., then

$$Y_j^{up} = E^{\mathcal{F}_j} Z_{\mathcal{J}} + E^{\mathcal{F}_j} \sum_{i=j}^{\mathcal{J}-1} [Z_i - E^{\mathcal{F}_i} Y_{i+1}]^+$$

is an upper approximation for  $Y_j^*$ , that is

$$Y_j \leq Y_j^* \leq Y_j^{up}, \quad j = 0, \dots, \mathcal{J}.$$

# Riesz upper bounds

## Proposition

- *Monotonicity*

$$\tilde{Y}_i \geq Y_i \longrightarrow \tilde{Y}_i^{up} \leq Y_i^{up}$$

- *Locality*

Let  $\{Y_i^\alpha, \alpha \in I_i\}$  be a family of local lower bounds at  $i$ , then

$$Y_j^{\alpha, up} = E^{\mathcal{F}_j} Z_{\mathcal{J}} + E^{\mathcal{F}_j} \sum_{i=j}^{\mathcal{J}-1} [Z_i - \max_{\alpha \in I_{i+1}} E^{\mathcal{F}_i} Y_{i+1}^\alpha]^+$$

is an upper bound.

# Doob-Meyer Martingale

For any martingale  $M_{T_j}$ , starting at  $M_0 = 0$ ,

$$Y_0^{up}(M) := E^{\mathcal{F}_0} \left[ \max_{0 \leq j \leq \mathcal{J}} (Z_j - M_j) \right]$$

is an **upper bound** for  $Y_0^*$ .

Exact value  $Y_0^*$  is attained at the martingale part  $M^*$  of the Snell envelope:

$$Y_j^* = Y_0^* + M_j^* - A_j^*,$$

where  $M_0^* = A_0^* = 0$  and  $A_j^*$  is  $\mathcal{F}_{j-1}$  measurable.



# Doob-Meyer Martingale

Let  $Y_j$  be an approximation for the Snell envelope  $Y_j^*$  with the Doob decomposition

$$Y_j = Y_0 + M_j - A_j,$$

where  $M_0 = A_0 = 0$  and  $A_j$  is  $\mathcal{F}_{j-1}$  measurable.

It then holds:

$$M_{j+1} - M_j = Y_{j+1} - \mathbb{E}^{\mathcal{F}_j}[Y_{j+1}]$$

## Observation

*The martingale  $M$  can be used to obtain an upper bound  $Y_0^{up}$  via*

$$Y_0^{up} = \mathbb{E} \left[ \max_{j=0, \dots, \mathcal{J}} (Z_j - M_j) \right].$$

# Andersen-Broadie Approach

Doob martingale corresponding to  $Y$ :

$$M_j = \sum_{i=1}^j (Y_i - \mathbb{E}_{\mathcal{F}_{i-1}} [Y_i]), \quad j = 0, \dots, \mathcal{J}.$$

Estimate the conditional expectations by Monte Carlo to get

$$M_j^k = \sum_{i=1}^j \left( Y_i - \frac{1}{k} \sum_{l=1}^k \xi_i^{(l)} \right), \quad k \in \mathbb{N},$$

where, conditionally on  $\mathcal{F}$ , all random variables  $\xi_i^{(l)}$ ,  $l = 1, \dots, k$ ,  $i = 1, \dots, \mathcal{J}$ , are independent and

$$\mathbb{E}_{\mathcal{F}} [\xi_i^{(l)}] = \mathbb{E}_{\mathcal{F}_{i-1}} [\xi_i^{(l)}] = \mathbb{E}_{\mathcal{F}_{i-1}} [\xi_i^{(1)}] = \mathbb{E}_{\mathcal{F}_{i-1}} [Y_i].$$

# Andersen-Broadie Approach

Fix some natural numbers  $N$  and  $K$ , and consider the estimator:

$$\begin{aligned} Y^{N,K} &= \frac{1}{N} \sum_{n=1}^N \max_{j=0,\dots,\mathcal{J}} (Z_j^{(n)} - M_j^{K,(n)}) \\ &=: \frac{1}{N} \sum_{n=1}^N \mathcal{Z}^{(n)}(M^K) \end{aligned}$$

based on a set of trajectories

$$\left\{ (Z_j^{(n)}, M_j^{K,(n)}), n = 1, \dots, N, j = 0, \dots, \mathcal{J} \right\}$$

of the vector process  $(Z, M^K)$ .

# Andersen-Broadie Approach

Denote

$$Y_0^{up} := E \left[ \max_{j=0, \dots, \mathcal{J}} (Z_j - M_j) \right] = E[\mathcal{Z}(M)]$$

## Observation

$$\begin{aligned} E \left[ (M_j^k - M_j)^2 \right] &= E \left[ \sum_{i=1}^j \left( E_{\mathcal{F}_{i-1}} [Y_i] - \frac{1}{k} \sum_{l=1}^k \xi_i^{(l)} \right) \right]^2 \\ &= \frac{1}{k} \sum_{i=1}^{\mathcal{J}} E \left[ E_{\mathcal{F}_{i-1}} [Y_i] - \xi_i^{(1)} \right]^2 \leq \frac{1}{k} \sum_{i=1}^{\mathcal{J}} E[Y_i^2] = O(1/k), \end{aligned}$$

*provided  $E[Y_j^2] < \infty$  for  $j = 0, \dots, \mathcal{J}$ .*

# Andersen-Broadie Approach

Then it holds

$$\begin{aligned} \mathbb{E} \left[ Y^{N,K} - Y_0^{up} \right]^2 &\leq N^{-1} \text{Var}(\mathcal{Z}(M^K)) + CK^{-\beta} \\ &=: N^{-1} v_K + CK^{-1}, \quad K \rightarrow \infty. \end{aligned}$$

In order to get  $\sqrt{\mathbb{E} [Y^{N,K} - Y(M)]^2} \leq \varepsilon$ , we may take

$$K = \left\lceil \frac{2C}{\varepsilon^2} \right\rceil,$$

and then

$$N = \left\lceil \frac{2v_K}{\varepsilon^2} \right\rceil$$

with  $\lceil x \rceil$  denoting the first integer which is larger than or equal to  $x$ .

# Andersen-Broadie Approach

If  $v_K$  is non-increasing, then, given an accuracy  $\varepsilon$ , the complexity is, up to a constant,

$$C^{N,K}(\varepsilon) := NK \lesssim \frac{V\left[\frac{2C}{\varepsilon^{2/\beta}}\right]}{\varepsilon^4}.$$

## Observation

If  $\text{Var}(\mathcal{Z}(M)) = 0$  (e.g.,  $M = M^*$ ) we have

$$v_K = \text{Var}(\mathcal{Z}(M^K)) \leq E(\mathcal{Z}(M^K) - \mathcal{Z}(M))^2 \leq BK^{-1},$$

and as a result

$$C^{N,K}(\varepsilon) \lesssim \frac{B}{2C} \frac{1}{\varepsilon^2}.$$

# Multilevel Approach

Let  $L \in \mathbb{N}$  and  $\mathbf{k} = (k_0, \dots, k_L)$  with  $1 \leq k_0 < k_1 < \dots < k_L$

$$\begin{aligned} Y(M^{k_L}) &= Y(M^{k_0}) + \sum_{l=1}^L [Y(M^{k_l}) - Y(M^{k_{l-1}})] \\ &= E[\mathcal{Z}(M^{k_0})] + \sum_{l=1}^L E[\mathcal{Z}(M^{k_l}) - \mathcal{Z}(M^{k_{l-1}})] \end{aligned}$$

with

$$Y(M^k) := E \left[ \max_{j=0, \dots, \mathcal{J}} (Z_j - M_j^k) \right] = E[\mathcal{Z}(M^k)]$$

# Multilevel Algorithm

- Fix a sequence  $\mathbf{n} = (n_0, \dots, n_L) \in \mathbb{N}^L$  with  $1 \leq n_0 < \dots < n_L$
- Simulate the initial set of trajectories:

$$\left\{ \left( Z_j^{(i)}, M_j^{k_0, (i)} \right), \quad i = 1, \dots, n_0, \quad j = 0, \dots, \mathcal{J} \right\}$$

of the vector process  $(Z, M^{k_0})$

- For each level  $l = 1, \dots, L$ , generate independently a set of trajectories:

$$\left\{ \left( Z_j^{(i)}, M_j^{k_{l-1}, (i)}, M_j^{k_l, (i)} \right), \quad i = 1, \dots, n_l, \quad j = 0, \dots, \mathcal{J} \right\}$$

of the vector process  $(Z, M^{k_{l-1}}, M^{k_l})$ .



# Multilevel Algorithm

Consider the approximation

$$\gamma^{n,k} := \frac{1}{n_0} \sum_{i=1}^{n_0} \mathcal{Z}^{(i)}(M^{k_0}) + \sum_{l=1}^L \frac{1}{n_l} \sum_{i=1}^{n_l} \left[ \mathcal{Z}^{(i)}(M^{k_l}) - \mathcal{Z}^{(i)}(M^{k_{l-1}}) \right]$$

with

$$\mathcal{Z}^{(i)}(M^k) := \max_{j=0,\dots,\mathcal{J}} \left( Z_j^{(i)} - M_j^{k,(i)} \right), \quad i = 1, \dots, n_l \quad k \in \mathbb{N}.$$

# Multilevel Algorithm

## Theorem

Let  $k_l = k_0 \kappa^l$ ,  $l = 0, \dots, L$ , for some  $k_0 \in \mathbb{N}$  and  $\kappa > 1$ . Set

$$L = \left\lceil -(\gamma \ln \kappa)^{-1} \ln \frac{\sqrt{k_0} \varepsilon}{C\sqrt{2}} \right\rceil$$

Let

$$n_l = \left\lceil 2\varepsilon^{-2}(L+1)k_0^{-1}\kappa^{-l} \right\rceil,$$

Then

$$\sqrt{\mathbb{E}[Y^{\mathbf{n}, \mathbf{k}} - Y(M)]^2} \leq \varepsilon,$$

while the *computational complexity* of the estimator  $Y^{\mathbf{n}, \mathbf{k}}$  is of order

$$\varepsilon^{-2} \ln^2 \varepsilon.$$



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